# On Distributions of Order Statistics from Nonidentical Discrete Variables

M. Güngör Department of Econometrics University of Inonu, Malatya, Turkey mgungor44@gmail.com

### **Abstract**

In this study, the distributions of rth order statistic of innid discrete random variables are obtained. In addition, the distributions are also expressed in the form of an integral. Then, the results related to pf and df of minimum and maximum order statistics of innid discrete random variables are given.

Keywords: Order statistics, Discrete random variable, Probability function,

Distribution function, Permanent.

MSC 2010: 62G30, 62E15.

#### 1. Introduction

The joint probability density function (pdf) and marginal pdf of order statistics of independent but not necessarily identically distributed(innid) random variables was derived by Vaughan and Venables (1972) by means of permanents. In addition, Balakrishnan (2007), and Bapat and Beg (1989) obtained the joint pdf and distribution function(df) of order statistics of *innid* random variables by means of permanents. In the first of two papers, Balasubramanian et al. (1991) obtained the distribution of single order statistic in terms of distribution functions of the minimum and maximum order statistics of some subsets of  $\{X_1, X_2, ..., X_n\}$ where  $X_i$ 's are *innid* random variables. Later, Balasubramanian et al. (1996) generalized their previous results (1991) to the case of the joint distribution function of several order statistics. Recurrence relationships among the distribution functions of order statistics arising from innid random variables were obtained by Cao and West (1997). Using multinomial arguments, the pdf of  $X_{rn+1}$  $(1 \le r \le n+1)$  was obtained by Childs and Balakrishnan (2006) by adding another independent random variable to the original n variables  $X_1, X_2, ..., X_n$ . Also, Balasubramanian et al. (1994) established the identities satisfied by distributions of order statistics from non-independent non-identical variables through operator methods based on the difference and differential operators. In a paper published in 1991, Beg obtained several recurrence relations and identities for product moments of order statistics of *innid* random variables using permanents. Recently, Cramer et al. (2009) derived the expressions for the distribution and density functions by Ryser's method and the distribution of maxima and minima based on permanents.

A multivariate generalization of classical order statistics for random samples from a continuous multivariate distribution was defined by Corley (1984). Guilbaud (1982) expressed the probability of the functions of innid random vectors as a linear combination of probabilities of the functions of independent and identically distributed(iid) random vectors and thus also for order statistics of random variables. Expressions for generalized joint densities of order statistics of iid random variables in terms of Radon-Nikodym derivatives with respect to product measures based on df were derived by Goldie and Maller (1999).

Several identities and recurrence relations for pdf and df of order statistics of iid random variables were established by numerous authors including Arnold et al. (1992), Balasubramanian and Beg (2003), David (1981), and Reiss (1989). Furthermore, Arnold et al. (1992), David (1981), Gan and Bain (1995), and Khatri (1962) obtained the probability function(pf) and df of order statistics of iid random variables from a discrete parent. Balakrishnan (1986) showed that several relations and identities that have been derived for order statistics from continuous distributions also hold for the discrete case. In a paper published in 1986, Nagaraja explored the behavior of higher order conditional probabilities of order statistics in a attempt to understand the structure of discrete order statistics. Later, Nagaraja (1992) considered some results on order statistics of a random sample taken from a discrete population.

In general, the distribution theory for order statistics is complicated when the discrete random variables are *innid*. In this study, the distributions of order statistics of *innid* discrete random variables are obtained.

As far as we know, these approaches have not been considered in the framework of order statistics from *innid* discrete random variables.

From now on, the subscripts and superscripts are defined in the first place in which they are used and these definitions will be valid unless they are redefined.

If  $a_1, a_2,...$  are defined as column vectors, then the matrix obtained by taking  $m_1$  copies of  $a_1$ ,  $m_2$  copies of  $a_2$ ,... can be denoted as

$$\begin{bmatrix} a_1 & a_2 & \dots \\ {}^{m_1} & {}^{m_2} & \end{bmatrix}$$

and *perA* denotes the permanent of a square matrix A, which is defined as similar to determinant except that all terms in the expansion have a positive sign.

Let  $X_1, X_2, ..., X_n$  be *innid* discrete random variables and  $X_{1:n} \le X_{2:n} \le ... \le X_{n:n}$  be the order statistics obtained by arranging the n  $X_i$ 's (i = 1, 2, ..., n) in increasing order of magnitude. Let  $F_i$  and  $f_i$  be df and pdf of  $X_i$ , respectively.

The paper is organized as follows. In section 2, we give the theorems concerning pf and df of order statistics of innid discrete random variables. In the last section, some results related to pf and df will be given.

## 2. Probability and distribution functions

In this section, the theorems related to pf and df of  $X_{rn}$  will be given. We will now express the following theorem for the pf of rth order statistic of innid discrete random variables.

## Theorem 2.1.

$$f_{r,n}(x) = \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(k+1+m)!(n-m-r)!} per[F(x-) \quad f(x) \quad 1-F(x)], \tag{1}$$

where  $F(x-) = (F_1(x-), F_2(x-), ..., F_n(x-))', \quad f(x) = (f_1(x), f_2(x), ..., f_n(x))',$ 1-  $F(x) = (I-F_1(x), I-F_2(x), ..., I-F_n(x))'$  are column vectors,  $F_i(x-) = P(X_i < x)$  and x = 0, 1, 2, ...

## **Proof.** Consider the event

$${X_{rn} = x}, r = 1,2,...,n.$$

The above event can be realized mutually exclusive ways as follows: r-1-k observations are less than x, k+1+m observations are equal to x and n-m-r observations are exceed x with respective probabilities  $F_i(x-)$ ,  $f_i(x)$  and  $1-F_i(x)$  (k=0,1,...,r-1 and m=0,1,...,n-r). Therefore, the pf of the above event can be written as

$$f_{r:n}(x) = P\{X_{r:n} = x\}.$$
 (2)

(2) can be expressed as (1). Thus, the proof is completed.

In Theorem 2.1., the pf of rth order statistic from innid discrete random variables is given by using permanent.

Consider (1). Using expansion of permanent in (1), the following theorem can be expressed .

## Theorem 2.2.

$$f_{r:n}(x) = \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(k+1+m)!(n-m-r)!} \sum_{p} \left( \prod_{l_1=1}^{r-1-k} F_{i_{l_1}}(x-) \right) \left( \prod_{l_2=r-k}^{r+m} f_{i_{l_2}}(x) \right) \cdot \prod_{l_3=r+m+1}^{n} [1 - F_{i_{l_3}}(x)],$$
(3)

where  $\sum_{p}$  denotes the sum over all n! permutations of  $(i_1, i_2, ..., i_n)$  of (1, 2, ..., n).

Proof. (1) can also be expressed as

$$f_{r,n}(x) = \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(k+1+m)!(n-m-r)!} \sum_{n_{s_1}, n_{s_2}} per[F(x-)][s_1/\cdot)$$

$$\cdot per[f(x)][s_2/\cdot) per[1-F(x)][s_3/\cdot),$$

where  $\sum_{n_{s_1},n_{s_2}}$  denotes the sum over  $\bigcup_{l=1}^2 s_l$  for which  $s_v \cap s_v = \phi$  for  $v \neq v$ ,  $s = \bigcup_{l=1}^3 s_l$ ,  $s_1 = \{s_1^{(1)}, s_1^{(2)}, \dots, s_1^{(r-1-k)}\}$ ,  $s_2 = \{s_2^{(1)}, s_2^{(2)}, \dots, s_2^{(k+1+m)}\}$ ,  $s_3 = \{s_3^{(1)}, s_3^{(2)}, \dots, s_3^{(n-m-r)}\}$ ,  $s = \{1, 2, \dots, n\}$ ,  $s_l \neq \phi$  and  $n_{s_l}$  is the cardinality of  $s_l$ . Futhermore,  $A[s_l/v)$  is the matrix obtained from A by taking rows whose indices are in  $s_l \subset s$ . Using expansion of permanent in the above identity, we get (3). Thus, the proof is completed.

Theorem 2.2. can also be written in form of an integral as follows.

#### Theorem 2.3.

$$f_{r:n}(x) = \frac{1}{(r-1)!(n-r)!} \sum_{P} \int_{F_{i_r}(x-)}^{F_{i_r}(x)} \left( \prod_{l_1=1}^{r-1} v_{i_{l_1}} \right) \left( \prod_{l_2=r+1}^{n} [1-v_{i_{l_2}}] \right) dv_{i_r},$$
where  $v_{i_j} = [v_{i_r} - F_{i_r}(x-)] \frac{f_{i_j}(x)}{f_{i_j}(x)} + F_{i_j}(x-), \ j = 1, 2, ..., n.$  (4)

**Proof.** Consider (3). It can be written as

$$f_{r:n}(x) = \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(k+1+m)!(n-m-r)!} \sum_{p} \left( \prod_{j_1=1}^{r-1-k} F_{i_{j_1}}(x-) \right) \left( \prod_{j_2=r-k}^{r-1} f_{i_{j_2}}(x) \right) \cdot f_{i_r}(x) \left( \prod_{j_3=r+1}^{r+m} f_{i_{j_3}}(x) \right) \left( \prod_{j_4=r+m+1}^{n} [1-F_{i_{j_4}}(x)] \right) \int_{0}^{1} \frac{(k+1+m)!}{k!m!} y^k (1-y)^m dy.$$
 (5)

Also, (5) can be clearly written as

$$f_{rn}(x) = \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(n-m-r)!k!m!} \sum_{P=0}^{1} \left( \prod_{j_1=1}^{r-1-k} F_{i_{j_1}}(x-1) \right) \left( \prod_{j_2=r-k}^{r-1-k} y f_{i_{j_2}}(x) \right) \cdot f_{i_r}(x) \left( \prod_{j_3=r+1}^{r+m} (1-y) f_{i_{j_3}}(x) \right) \left( \prod_{j_4=r+m+1}^{n} [1-F_{i_{j_4}}(x)] \right) dy.$$
 (6)

In (6), if  $v_{i_i} = y f_{i_i}(x) + F_{i_i}(x-)$ , the following identity is obtained

$$f_{r,n}(x) = \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)! k! m! (n-m-r)!} \sum_{P} \int_{F_{i_r}(x-1)}^{F_{i_r}(x)} \left( \prod_{j_1=1}^{r-1-k} F_{i_{j_1}}(x-1) \right) dv_{i_r} dv_{i_r} dv_{i_1} dv_{i_2} dv_{i_3} dv_{i_4} dv_{i_5} dv_{i_5} dv_{i_6} dv_{i_7} dv_{i_5} dv_{i_7} dv_{i_$$

Now consider the following identity

$$\sum_{k=0}^{r-1} \frac{1}{(r-1-k)!k!} \sum_{P} \left( \prod_{j_1=1}^{r-1-k} A_{i_{j_1}}(x) \right) \prod_{j_2=r-k}^{r-1} B_{i_{j_2}}(x) = \frac{1}{(r-1)!} \sum_{P} \prod_{l=1}^{r-1} [A_{i_l}(x) + B_{i_l}(x)]. \tag{8}$$

Using (8) in (7), we get (4). Thus, the proof is completed.

We will now express the following theorem to obtain the df of rth order statistic from innid discrete random variables.

## Theorem 2.4.

$$F_{r:n}(x) = \sum_{z=0}^{x} \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(k+1+m)!(n-m-r)!} per[F(z-) \quad f(z) \quad 1-F(z)].$$
 (9)

Proof. It can be written

$$F_{r:n}(x) = \sum_{n=0}^{x} f_{r:n}(z)$$
 (10)

and using (1) in (10), (9) is obtained.

Using expansion of permanent in the above theorem, we will establish the following theorem.

## Theorem 2.5.

$$F_{r:n}(x) = \sum_{z=0}^{x} \sum_{k=0}^{r-1} \sum_{m=0}^{n-r} \frac{1}{(r-1-k)!(k+1+m)!(n-m-r)!} \sum_{p} \left( \prod_{l_1=1}^{r-1-k} F_{i_{l_1}}(z-) \right) \left( \prod_{l_2=r-k}^{r+m} f_{i_{l_2}}(z) \right) \cdot \prod_{l_3=r+m+1}^{n} [1 - F_{i_{l_3}}(z)].$$
(11)

**Proof.** Using (3) in (10), (11) is obtained.

(9) can also be written in form of an integral as follows.

## Theorem 2.6.

$$F_{r,n}(x) = \frac{1}{(r-1)!(n-r)!} \sum_{P} \int_{0}^{F_{i_r}(x)} \left( \prod_{l_1=1}^{r-1} v_{i_{l_1}} \right) \left( \prod_{l_2=r+1}^{n} [1-v_{i_{l_2}}] \right) dv_{i_r}.$$
 (12)

**Proof.** Using (4) in (10), (12) is obtained.

# 3. Results for probability and distribution functions

In this section, the results related to pf and df of  $X_{r:n}$  will be given. We will now express the following result for pf of minimum order statistic of innid discrete random variables.

#### Result 3.1.

$$f_{1:n}(x) = \sum_{m=0}^{n-1} \frac{1}{(1+m)!(n-m-1)!} per[f(x) \quad 1-F(x)]$$

$$= \sum_{m=0}^{n-1} \frac{1}{(1+m)!(n-m-1)!} \sum_{P} \left( \prod_{l_2=1}^{1+m} f_{i_{l_2}}(x) \right) \prod_{l_3=m+2}^{n} [1-F_{i_{l_3}}(x)]$$

$$= \frac{1}{(n-1)!} \sum_{P} \int_{F_{i_1}(x-1)}^{F_{i_1}(x)} \left( \prod_{l_2=2}^{n} [1-v_{i_{l_2}}] \right) dv_{i_1}.$$
(13)

**Proof.** In (1), (3) and (4), if r = 1, (13) is obtained.

We will express the following result for pf of maximum order statistic of *innid* discrete random variables.

#### Result 3.2.

$$f_{n:n}(x) = \sum_{k=0}^{n-1} \frac{1}{(n-1-k)!(k+1)!} per[F(x-) \quad f(x)]$$

$$= \sum_{k=0}^{n-1} \frac{1}{(n-1-k)!(k+1)!} \sum_{P} \left( \prod_{l_1=1}^{n-1-k} F_{i_{l_1}}(x-) \right) \prod_{l_2=n-k}^{n} f_{i_{l_2}}(x)$$

$$= \frac{1}{(n-1)!} \sum_{P} \sum_{F_{i_n}(x-)}^{F_{i_n}(x)} \left( \prod_{l_1=1}^{n-1} v_{i_{l_1}} \right) dv_{i_n}.$$
(14)

**Proof.** In (1), (3) and (4), if r = n, (14) is obtained.

In the following result, we will express df of minimum order statistic of *innid* discrete random variables.

#### Result 3.3.

$$F_{1:n}(x) = \sum_{z=0}^{x} \sum_{m=0}^{n-1} \frac{1}{(1+m)!(n-m-1)!} per[f(z) \quad 1 - F(z)]$$

$$= \sum_{z=0}^{x} \sum_{m=0}^{n-1} \frac{1}{(1+m)!(n-m-1)!} \sum_{P} \left( \prod_{l_2=1}^{1+m} f_{i_{l_2}}(z) \right) \prod_{l_3=m+2}^{n} [1 - F_{i_{l_3}}(z)]$$

$$= \frac{1}{(n-1)!} \sum_{P} \int_{0}^{F_{i_1}(x)} \left( \prod_{l_2=2}^{n} [1 - v_{i_{l_2}}] \right) dv_{i_1}.$$
(15)

**Proof.** In (9), (11) and (12), if r = 1, (15) is obtained.

We will express the following result for df of maximum order statistic of *innid* discrete random variables.

#### Result 3.4.

$$F_{n:n}(x) = \sum_{z=0}^{x} \sum_{k=0}^{n-1} \frac{1}{(n-1-k)!(k+1)!} per[F(z-) f(z)]$$

$$= \sum_{z=0}^{x} \sum_{k=0}^{n-1} \frac{1}{(n-1-k)!(k+1)!} \sum_{P} \left( \prod_{l_1=1}^{n-1-k} F_{i_{l_1}}(z-) \right) \prod_{l_2=n-k}^{n} f_{i_{l_2}}(z)$$

$$= \frac{1}{(n-1)!} \sum_{P} \sum_{0}^{F_{i_n}(x)} \left( \prod_{l_1=1}^{n-1} v_{i_{l_1}} \right) dv_{i_n}.$$
(16)

**Proof.** In (9), (11) and (12), if r = n, (16) is obtained.

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