The Kumaraswamy- Laplace Distribution

Manal Mohamed Nassar Department of Mathematics Ain Shams University, Cairo, Egypt manal.m.nassar@gmail.com

Abstract

A generalized Laplace distribution using the Kumaraswamy distribution is introduced. Different structural properties of this new distribution are derived, including the moments, and the moment generating function. We discuss maximum likelihood estimation of the model parameters and obtain the observed and expected information matrix. A real data set is used to compare the new model with widely known distributions.

Keywords: Kumaraswamy- Laplace distribution, Maximum likelihood estimation, Akaike Information Criterion and Bayesian Information Criterion.

1. Introduction

For many years, addition of a positive real valued parameter \(\Bar{\pi} \) to any cumulative distribution function \Box cdf \Box G through exponentiation produces another cdf labeled G^a is always richer and more flexible for data modeling. Mudhokar and Srivastava (1993) successfully applied the exponentiatedWeibull (EW) to analyze bathtub failure data. Gupta and Kundu (2001) concentrated on the study of the exponentiated exponential and found out that it describes situations better than the Weibull or gamma but at least in certain circumstances Exponentiated Exponential might work better than Weibull or gamma. Nassar and Eissa (2003&2004) studied some properties and Bayesian estimates of the exponentiated Weibull (EW) distribution. Nadarajah (2004) applied the exponentiated Gumbel distribution to climate data with improved desired results. Later, the Beta distribution, which is one of the most basic distributions supported on finite range (0, 1), has been used widely in both practical and theoretical generalizations aspects of statistics (see Nadarajah and Kotz (2004&2005), Kong et al. (2007), Akinsete et al. (2008) Pescim et al. (2010), Souza et al. (2010), Nassar and Elmasry (2012) and Nassar and Nada (2011 & 2012 & 2013). An alternative distribution like the beta distribution, which is easier to work with, is the Kumaraswamy distribution proposed by Kumaraswamy (1980). This distribution has a simple form, where the probability density function (pdf) and the cumulative distribution function (cdf) are given by respectively:

$$f(x) = abx^{a-1}(1 - x^a)^{b-1} \quad ; 0 < x < 1$$
 (1)

$$F(x) = [1 - (1 - x^a)^b]; 0 < x < 1$$
(2)

where a>0, b>0 are the shape parameters. The Kumaraswamy distribution is like the beta distribution in many ways, for example, Kumaraswamy's densities are also unimodal, uniantimodal, increasing, decreasing or constant depending in the same way as the beta distribution on the values of its parameters. In addition, one can easily show that the Kumaraswamy distribution has the same basic shape properties of the beta distribution. But, because the cdf of the Kumaraswamy distribution has a simple closed form, it has received much attention in simulating hydrological data and related areas. For more

detailed properties of the Kumaraswamy distribution, see Jones (2008). Cordeiro and Lemonte (2011) introduced the beta Laplace distribution, along with the paper by Xiaohu et al. (2011) where they discussed the Kumaraswamy Binomial distribution model showing that it performs better than Beta-Binomial model in some occasions. De Santana et al. (2012) introduced the Kumaraswamy-log-logistic distribution which contains several important distributions as sub-models such as the Log- Logistic, exponentiated Log-Logistic and Burr distributions. Bidram and Nekoukhou (2013) introduced double-bounded Kumaraswamy power series class of distribution, where this new class obtained by compounding the Kumaraswamy distribution and the family of power series distribution. Nadarajah and Eljabri (2013) discussed the Kumaraswamy general Pareto distribution. Akinseteet al. (2014) discussed the Kumaraswamy-geometric distribution.

In this article we propose a new model, the so-called Kumaraswamy Laplace (KL) distribution. Thearticle is outlined as follows. In Section 2 we introduce the KL distribution and provide plots of density function. We derive the r-th moment and hence obtain the expected value, variance and the moment generating function in Section 3. In Section 4, we deduce the maximum likelihood estimates of the parameters, and derive the information matrix. In Section 5, we provide the Renyi and Shannon entropies. The importance of the (KL) distribution through the analysis of a real data set is illustrated in Section 7.

2. The Kumaraswamy Laplace distribution

One of the earliest distributions in probability theory was introduced by Pierre-Simon Laplace in 1774. A random variable Z has the Laplace distribution with location parameter μ and scale parameter $\sigma>0$, if its probability density function (pdf) is given by

$$g(z) = \frac{1}{2\alpha} e^{\left[-\frac{|z-\mu|}{\alpha}\right]}, \quad -\infty < z < \infty$$
 (3)

The cumulative distribution function becomes

$$G(z) = \begin{cases} \frac{1}{2} e^{\left[\frac{z-\mu}{\alpha}\right]} z < \mu \\ 1 - \frac{1}{2} e^{\left[-\frac{z-\mu}{\alpha}\right]} z \ge \mu \end{cases}$$
 (4)

The cumulative distribution function of the Kumaraswamy Laplace (KL) distribution can be written as

$$F(z) = \begin{cases} 1 - (1 - (\frac{1}{2}e^{\frac{z-\mu}{\alpha}})^a)^b & z < \mu \\ 1 - (1 - (1 - \frac{1}{2}e^{\frac{z-\mu}{\alpha}})^a)^b & z \ge \mu \end{cases}$$
 (5)

The density function corresponding to (5) is given by

$$f(z) = \begin{cases} \frac{2^{-a}ab(e^{\frac{-\mu+z}{\alpha}})^{a}(1-2^{-a}(e^{\frac{-\mu+z}{\alpha}})^{a})^{-1+b}}{\alpha} & z < \mu \\ \frac{abe^{-\frac{-\mu+z}{\alpha}}(1-\frac{1}{2}e^{-\frac{-\mu+z}{\alpha}})^{-1+a}(1-(1-\frac{1}{2}e^{-\frac{-\mu+z}{\alpha}})^{a})^{-1+b}}{2\alpha} & z \ge \mu \end{cases}$$
(6)

If we put a=b=1 in Equation (6), introduces to the standard Laplace density function given by Equation (3). If Z follows (6), we can say that $Z \sim KL(a, b, \alpha, \mu)$. By taking the transformation

$$X = \frac{Z - \mu}{\alpha}$$

Equation (6) can be written in a simpler form as

$$f(x) = \begin{cases} 2^{-a}abe^{ax}(1 - 2^{-a}e^{ax})^{-1+b} & x < 0\\ \frac{1}{2}abe^{-x}(1 - \frac{1}{2}e^{-x})^{-1+a}(1 - (1 - \frac{1}{2}e^{-x})^{a})^{-1+b} & x \ge 0 \end{cases}$$
(7)

The graphs in Figure 1 outline different shapes of the density functions for various values of the parameters.

3. Moment generating function and moments

The most important features of the KL distribution can be studied through moments. We know the r-th moment for a random variable X following the distribution given by Equation (7) can be written as

$$E(x^{r}) = \int_{-\infty}^{0} \frac{ab}{2^{a}} e^{ax} (1 - 2^{-a}e^{ax})^{b-1} x^{r} dx + \int_{0}^{\infty} \frac{ab}{2} e^{-x} (1 - \frac{1}{2}e^{-x})^{a-1} (1 - (1 - \frac{1}{2}e^{-x})^{a})^{b-1} x^{r} dx$$
(8)

Considering a, b are real non-integer, then we have the power series expansion

$$(1-v)^c = \sum_{l=0}^{\infty} (-1)^l {c \choose l} v^l, |v| < 1$$

By applying this expansion for equation (8), we get the general form of the r-th moments about zero as:

$$E(x^r) = m \int_{-\infty}^0 x^r e^{a(l+1)x} dx + k \int_0^\infty x^r e^{-x(l+1)} dx$$
 (9)

Where

$$m = \frac{ab}{2a} \sum_{l=0}^{\infty} 2^{-al} (-1)^{l} {b-1 \choose l}$$
 (10)

And

$$k = \frac{ab}{2} \sum_{j=0}^{\infty} \sum_{i=0}^{\infty} (-1)^{i+j} {b-1 \choose i} {a(j+1)-1 \choose i} 2^{-i}$$
(11)

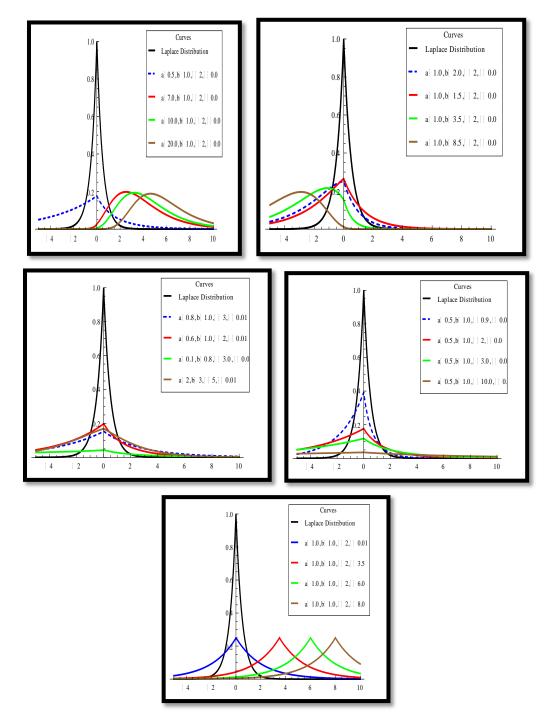


Fig 1. Plots of the pdf of KL for various values of parameters.

Then

$$E(x^r) = m \frac{(-1)^r}{[a(l+1)]^{1+r}} \Gamma(1+r) + k \frac{1}{(i+1)^r} \Gamma(1+r)$$
(12)

where a(l+1) > 0, Re(r) > -1, and Re(i) > -1. From Equation (12), we can get the first moment (Expected value) and the variance of (KL) distribution, which are given, respectively, as follows

$$E(x) = \frac{k}{(1+i)^2} - \frac{m}{(a^2(1+l)^2)}$$
 (13)

where a(l+1) > 0

$$Var(x) = \frac{2m}{a^2(l+1)^3} + \frac{2k}{(l+1)^3} - \left[\left(\frac{m}{a^2(l+1)^2} - \frac{k}{(l+1)^2}\right)^2\right]$$
(14)

Where m and k are defined before in Equations (10) and (11).

The moment generating function can be written from Equation (7) as

$$E[e^{tx}] = m \int_{-\infty}^{0} e^{x(t+a(t+1))} dx + k \int_{0}^{\infty} e^{x(t-(t+1))} dx$$
to obtain
$$E[e^{tx}] = \frac{m}{t+a(t+1)} + \frac{k}{1+t-t}$$
(15)

$$t+a(l+1)$$
 $1+l-t$

Where a(1+l) + t > 0, i-t > -1 and m and k are defined in Equations (10) and (11).

4. Estimation

Let $Z_1, Z_2, ..., Z_n$ be a random sample from KL distribution given by Equation (6). The parameters of the KL distribution are estimated by the maximum likelihood method and we consider $\theta = (a, b, \alpha, \mu)^T$ be the unknown parameter vector. The Log-Likelihood function for θ is given by

$$l(\theta) = \sum_{i=1}^{n} \delta_{i} Log(\frac{ab}{2^{\alpha}\alpha}) + \sum_{i=1}^{n} \delta_{i} \left[a\left(\frac{z_{i}-\mu}{\alpha}\right) + (b-1)\log\left[1 - \left(\frac{1}{2}e^{\left(\frac{z_{i}-\mu}{\alpha}\right)^{a}}\right)\right] + \\ + \sum_{i=1}^{n} (1 - \delta_{i}) Log(\frac{ab}{2^{\alpha}}) + \sum_{i=1}^{n} (1 - \delta_{i})\left(\left(\frac{\mu-z_{i}}{\alpha}\right) + (a-1)\log\left[1 - \frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}}\right]\right] \\ + (b-1)\log(1 - \left(1 - \frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}}\right)^{a}$$

$$\text{where } \delta_{i} = \begin{cases} 1; & \text{if } Z_{i} < \mu; \\ 0; & \text{if } Z_{i} \ge \mu \end{cases}, & \text{for } i = 1, \dots, n.$$

The associated score function is given by $u(\theta) = (u_a, u_b, u_\alpha, u_\mu)^T$ where

$$\begin{split} u_{a} &= \sum_{i=1}^{n} \frac{1-\delta_{i}}{a} + \sum_{i=1}^{n} (1-\delta_{i}) \log(1-\frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}}) - \frac{(b-1)(1-\frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}})^{a} \log(1-\frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}})}{1-(1-\frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}})^{a}} + \\ &+ \sum_{i=1}^{n} \delta_{i} \frac{2^{a} \alpha (\frac{(2^{-a}b-2^{-a}ablog(2)}{\alpha})}{ab} + \\ &+ \sum_{i=1}^{n} \delta_{i} (\frac{(b-1)(2^{-a}(e^{\frac{-\mu+z_{i}}{\alpha}})^{a} \log(2) - 2^{-a}(e^{\frac{-\mu+z_{i}}{\alpha}})^{a} \log(e^{\frac{-\mu+z_{i}}{\alpha}})}{1-2^{-a}(e^{\frac{-\mu+z_{i}}{\alpha}})^{a}} + \frac{-\mu+z_{i}}{\alpha}) \\ &+ \sum_{i=1}^{n} \delta_{i} (\frac{1-\delta_{i}}{a}) (1-\delta_{i}) \left[\frac{1}{b} + \log(1-(1-\frac{1}{2}e^{\frac{\mu-z_{i}}{\alpha}})^{a})\right] + \sum_{i=1}^{n} \delta_{i} \left[\frac{1}{b} + \log(1-2^{-a}(e^{\frac{-\mu+z_{i}}{\alpha}})^{a})\right] \end{split}$$

$$u_{\mu} = \sum_{i=1}^{n} \left(\frac{1}{\alpha} - \frac{(\alpha - 1)e^{\frac{\mu - z_{i}}{\alpha}}}{2\left(1 - \frac{1}{2}e^{\frac{\mu - z_{i}}{\alpha}}\right)\alpha} + \frac{a(b - 1)e^{\frac{\mu - z_{i}}{\alpha}}\left(1 - \frac{1}{2}e^{\frac{\mu - z_{i}}{\alpha}}\right)^{-1 + \alpha}}{2\left(1 - \left(1 - \frac{1}{2}e^{\frac{\mu - z_{i}}{\alpha}}\right)^{a}\right)\alpha} \right) (1 - \delta_{i})$$

$$+ \sum_{i=1}^{n} \left(-\frac{a}{\alpha} + \frac{2^{-a}a(b - 1)(e^{\frac{-\mu + z_{i}}{\alpha}})^{a}}{(1 - 2^{-a}\left(e^{\frac{-\mu + z_{i}}{\alpha}}\right)^{a})\alpha} \right)$$

$$u_{\alpha} = \sum_{i=1}^{n} -\frac{1 - \delta_{i}}{\alpha}$$

$$+ \sum_{i=1}^{n} \left(-\frac{\mu - z_{i}}{\alpha^{2}} + \frac{(\alpha - 1)e^{\frac{\mu - z_{i}}{\alpha}}(\mu - z_{i})}{2\left(1 - \frac{1}{2}e^{\frac{\mu - z_{i}}{\alpha}}\right)\alpha^{2}} \right)$$

$$- \frac{a(b - 1)e^{\frac{\mu - z_{i}}{\alpha}}\left(1 - \frac{1}{2}e^{\frac{\mu - z_{i}}{\alpha}}\right)^{-1 + a}(\mu - z_{i})}{2(1 - \left(1 - \frac{1}{2}e^{\frac{\mu - z_{i}}{\alpha}}\right)^{a}\alpha^{2}} \right)$$

$$+ \sum_{i=1}^{n} -\frac{\delta_{i}}{\alpha}$$

$$+ \sum_{i=1}^{n} \left(-\frac{a(-\mu + z_{i})}{\alpha^{2}} + \frac{2^{-a}a(b - 1)\left(e^{\frac{-\mu + z_{i}}{\alpha}}\right)^{a}(-\mu + z_{i})}{(1 - 2^{-a}\left(e^{\frac{-\mu + z_{i}}{\alpha}}\right)^{a})\alpha^{2}} \right) \delta_{i}$$

The maximum likelihood estimation (MLE) of θ is $\hat{\theta}$ is obtained by setting the previous equations to zero $U(\hat{\theta}) = 0$, and solving them using Mathematica. The normal approximation of the MLE of θ can be used for constructing approximate confidence intervals and for testing hypotheses on the parameters a, b, α and μ . Now we derive the Fisher information matrix for interval estimation and hypotheses testing on the model parameters. The 4x4 Fisher information matrix is given by:

$$\sum(\theta) = -\begin{pmatrix} u_{aa}u_{ab}u_{a\alpha}u_{a\mu} \\ u_{ba}u_{bb}u_{b\alpha}u_{b\mu} \\ u_{\alpha a}u_{\alpha b}u_{\alpha \alpha}u_{\alpha \mu} \\ u_{\mu a}u_{\mu b}u_{\mu \alpha}u_{\mu \mu} \end{pmatrix}$$

Whose elements are given by the following

$$u_{\rm bb} = -\frac{1}{b^2} \left[\sum_{i=1}^n \delta_i + \sum_{i=1}^n (1 - \delta_i) \right] = -\frac{n}{b^2}.$$

The Kumaraswamy- Laplace Distribution

$$\begin{split} u_{aa} &= \sum_{i=1}^{n} (1-\delta_{i})[-\frac{1}{a^{2}} \\ &- (b-1)\{\frac{\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a} 2Log\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a}\right)} \left[\frac{\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a}}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a}\right)} \\ &+ 1\right]\}] + \sum_{i=1}^{n} \delta_{i} \left[\frac{-(b-1)\left(\frac{1}{2}e^{\frac{t_{i}-\mu}{\alpha}}\right)^{2a}}{\left(1-\left(\frac{1}{2}e^{\frac{t_{i}-\mu}{\alpha}}\right)^{a}}\right)^{2}} \left[Log(2) - Log\left(e^{\frac{t_{i}-\mu}{\alpha}}\right)\right]^{2} \\ &+ \frac{2Log(2)(b-1)\left(\frac{1}{2}e^{\frac{t_{i}-\mu}{\alpha}}\right)^{a}}{\left(1-\left(\frac{1}{2}e^{\frac{t_{i}-\mu}{\alpha}}\right)^{a}}\right)} \{1 + Log\left(e^{\frac{t_{i}-\mu}{\alpha}}\right) - \frac{Log\left(e^{\frac{t_{i}-\mu}{\alpha}}\right)}{Log(2)}\} - (\frac{1}{a^{2}}e^{\frac{t_{i}-\mu}{\alpha}}e^{\frac$$

$$\begin{split} u_{\alpha\alpha} &= \sum_{i=1}^{n} (1-\delta_{i}) \right] \left[\left(\frac{\mu - \mathsf{t}_{i}}{\alpha^{3}} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right) \left\{ 2e^{-\left(\frac{\mu - \mathsf{t}_{i}}{\alpha} \right)} + \left[-\frac{\alpha - 1}{\left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)} \right\} \right. \\ &\left. + \frac{e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} (\mu - \mathsf{t}_{i})}{4\alpha \left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)} \right\} \\ &\left. + \frac{a(b-1) \left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)^{a-1}}{\left(1 - \left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)^{a} \right)} \left[1 - \frac{ae^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)^{a-1} (\mu - \mathsf{t}_{i})}{4\alpha \left(1 - \left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)^{a} \right)} - \frac{(a-1)e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}}}{4\alpha \left(1 - \frac{1}{2} e^{\frac{\mu - \mathsf{t}_{i}}{\alpha}} \right)} \right] \\ &\left. + \frac{1}{\alpha^{2}} \right] + \sum_{i=1}^{n} \delta_{i} \left[\frac{1}{\alpha^{2}} + \frac{2a(e^{\frac{\mathsf{t}_{i} - \mu}{\alpha}})^{a} (\mathsf{t}_{i} - \mu)}{\alpha^{3} \left(1 - \left(\frac{1}{2} e^{\frac{\mathsf{t}_{i} - \mu}{\alpha}} \right)^{a} \right)} \left[\left(\left(1 - \left(\frac{1}{2} e^{\frac{\mathsf{t}_{i} - \mu}{\alpha}} \right)^{a} \right) \left(e^{\frac{\mathsf{t}_{i} - \mu}{\alpha}} \right)^{a} - 2^{a} (b-1) - \frac{2^{-(2a+1)}a(b-1)(e^{\frac{\mathsf{t}_{i} - \mu}{\alpha}})^{a} (\mathsf{t}_{i} - \mu)}{\alpha} - \frac{2^{-(a+1)}}{\alpha} a(b-1)(\mathsf{t}_{i} - \mu) \right] \end{split}$$

$$\begin{split} & \sum_{i=1}^{n} - (1-\delta_{i}) \left[\frac{\frac{e^{\frac{\mu-t_{i}}{\alpha}}}{4\alpha^{2}}}{\left[\frac{(a-1)}{\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)} \left\{ \frac{e^{\frac{\mu-t_{i}}{\alpha}}}{\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)} + 2 \right\} + \\ & \left\{ \frac{a(b-1)\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a-1}}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a}\right)} \left[ae^{\frac{\mu-t_{i}}{\alpha}} \frac{\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a-1}}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)^{a}\right)} + \frac{(a-1)e^{\frac{\mu-t_{i}}{\alpha}}}{\left(1-\frac{1}{2}e^{\frac{\mu-t_{i}}{\alpha}}\right)} - 2 \right] \right\} \right] + \\ & \sum_{i=1}^{n} \delta_{i} \left[\frac{-2^{-a}a^{2}(b-1)}{\alpha^{2}\left(1-\left(\frac{1}{2}e^{\frac{t_{i}-\mu}{\alpha}}\right)^{a}\right)^{2}} \left\{ 2^{-a}e^{\frac{t_{i}-\mu}{\alpha}} + 1 \right\} \right] \end{split}$$

$$u_{ab} = \sum_{i=1}^{n} -(1-\delta_{i}) Log(1-\frac{1}{2}e^{\frac{\mu-\mathsf{t}_{i}}{\alpha}}) [\frac{(1-\frac{1}{2}e^{\frac{\mu-\mathsf{t}_{i}}{\alpha}})^{a}}{1-(1-\frac{1}{2}e^{\frac{\mu-\mathsf{t}_{i}}{\alpha}})^{a}}] - \sum_{i=1}^{n} \delta_{i} \{ Log(\frac{1}{2}e^{\frac{\mathsf{t}_{i}-\mu}{\alpha}}) [\frac{(\frac{1}{2}e^{\frac{\mathsf{t}_{i}-\mu}{\alpha}})^{a}}{1-(\frac{1}{2}e^{\frac{\mathsf{t}_{i}-\mu}{\alpha}})^{a}}] \}$$

 $u_{\mu\mu} =$

$$\begin{split} u_{b\mu} &= \frac{2^{-a}a^2}{\alpha} [\sum_{i=1}^n (1-\delta_i) [\frac{2^{a-1}e^{\frac{\mu-\mathbf{t}_i}{\alpha}} \left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^{a-1}}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^a\right)} + \sum_{i=1}^n \delta_i \frac{(e^{\frac{\mathbf{t}_i-\mu}{\alpha}})^a}{\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^a}]] \\ u_{\alpha\mu} &= \sum_{i=1}^n -\frac{ae^{\frac{\mu-\mathbf{t}_i}{\alpha}} \left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^{a-1} (\mu-\mathbf{t}_i)}{2\left(1-((1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}})^a\right)\alpha^2} (1-\delta_i) + \sum_{i=1}^n \frac{a\left(\frac{1}{2}e^{\frac{\mathbf{t}_i-\mu}{\alpha}}\right)^a (\mathbf{t}_i-\mu)}{(1-\left(\frac{1}{2}e^{\frac{\mathbf{t}_i-\mu}{\alpha}}\right)^a\right)\alpha^2} \delta_i \\ u_{a\mu} &= \sum_{i=1}^n (1-\delta_i) [\frac{e^{\frac{\mu-\mathbf{t}_i}{\alpha}}}{2\alpha(1-\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^a)} [\frac{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^a\right)}{\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)} \\ &+ a(b-1)\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^{a-1} \left\{\frac{1}{a}\right. \\ &+ \log\left(\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)\left[\frac{\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^a}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-\mathbf{t}_i}{\alpha}}\right)^a\right)} + 1\right]\right\} \\ &+ \sum_{i=1}^n \delta_i [\frac{1}{\alpha}[-1+2^{-a}\alpha(b-1)(e^{\frac{\mathbf{t}_i-\mu}{\alpha}})^a \{2^{-a}a(e^{\frac{\mathbf{t}_i-\mu}{\alpha}})^a Log\left(\frac{1}{2}e^{\frac{\mathbf{t}_i-\mu}{\alpha}}\right) + \frac{1}{\left(1-\left(\frac{1}{2}e^{\frac{\mathbf{t}_i-\mu}{\alpha}}\right)^a\right)} [1 + Log\left(\frac{1}{2}e^{\frac{\mathbf{t}_i-\mu}{\alpha}}\right)^a]\}]] \end{split}$$

$$\begin{split} u_{a\alpha} &= \sum_{i=1}^{n} (1-\delta_i) \frac{(\mu-t_i)(e^{\frac{\mu-t_i}{\alpha}})}{2\alpha^2} \Bigg[\frac{1}{\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)} - \frac{(b-1)\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^{a-1}}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^a\right)} \\ &- \frac{a(b-1)\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^{a-1}\log \left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^a\right)} \Bigg\} \\ &+ \frac{\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^a}{\left(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^a\right)} \Bigg\} \\ &+ \sum_{i=1}^n \delta_i \Big[-\frac{t_i-\mu}{\alpha^2} \\ &+ \frac{(b-1)\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right)^a(t_i-\mu)}{\left(1-\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right)^a\right)\alpha^2} \Big\{ 1-aLog(\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right) \\ &+ \frac{a\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right)^a}{2}Log(\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right)} \Big\} \Big] \\ u_{b\alpha} &= \sum_{i=1}^n (1-\delta_i)(-\frac{ae^{\frac{\mu-t_i}{\alpha}}(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}})^{a-1}(\mu-t_i)}{2(1-\left(1-\frac{1}{2}e^{\frac{\mu-t_i}{\alpha}}\right)^a)\alpha^2} \\ &+ \sum_{i=1}^n \delta_i(\frac{a\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right)^a(t_i-\mu)}{\left(1-\left(\frac{1}{2}e^{\frac{t_i-\mu}{\alpha}}\right)^a(t_i-\mu)}\right)} \Big\} \end{aligned}$$

Now, we derive the Expected information matrix, whose elements are given by the relation $I_{\theta_i\theta_j} = E\left[-\frac{\partial^2 l_n}{\partial \theta_i \partial \theta_i}\right]$; I, j =1, 2, 3, 4.

$$E[u_{bb}] = -$$

$$\left[\sum_{i=1}^{n}\sum_{j=0}^{\infty}\delta_{i}\binom{b-1}{j}\frac{2^{-a(j+1)}}{b(j+1)} - \frac{a}{2b}\sum_{i=1}^{n}\sum_{l=0}^{\infty}\sum_{k=0}^{\infty}\sum_{d=0}^{\infty}(1-\delta_{i})(-1)^{\binom{l}{l+k+d}}\frac{2^{-(l+d)}}{l+d+1}\right]$$

$$E[u_{ab}]$$
=-[-

$$\sum_{i=1}^{n} \sum_{n=1}^{\infty} \sum_{l=0}^{\infty} \sum_{j=0}^{\infty} \delta_{i} (-1)^{l+j} \frac{ab}{n} \binom{n}{l} \binom{b-2}{j} \frac{2^{l-a(j+1)}}{a[2+j]-l} + \frac{ab}{n} \binom{n}{l} \binom{b-2}{j} \frac{2^{l-a(j+1)}}{a[2+j]-l} + \frac{ab}{n} \binom{n}{l} \binom{n}$$

$$\sum_{i=1}^{n} \sum_{n=0}^{\infty} \sum_{o=0}^{\infty} \sum_{k=0}^{\infty} \sum_{d=0}^{\infty} (1-\delta_i) ab (-1)^{m+k+d} {2a-1 \choose m} {b-2 \choose k} {ak \choose d} \frac{2^{-(m+o+d+1)}}{(m+o+d+1)}$$

$$\begin{split} E \left[u_{b\mu} \right] &= - \left[\sum_{i=1}^{n} \sum_{j=0}^{\infty} \delta_{i} \frac{a^{2}b}{\alpha 2^{a(j+2)}} (-1)^{j} \binom{b-2}{j} \frac{1}{a(j+2)} + \sum_{i=1}^{n} \sum_{l=0}^{\infty} \sum_{d=0}^{\infty} \sum_{k=0}^{\infty} (1-\delta_{i}) \frac{a^{2}b}{\alpha} (-1)^{l+k+d} \binom{2a-2}{l} \binom{b-2}{j} \binom{ad}{k} \frac{2^{-(l+k+2)}}{l+k+2} \right] \end{split}$$

$$E[u_{\alpha\mu}] = -\left[\sum_{i=1}^{n} \sum_{j=0}^{\infty} \delta_{i}(-1)^{j} \frac{a^{2}b}{\alpha} \frac{2^{-a(j+2)}}{-(aj+2a)^{2}} \binom{b-2}{j} - \sum_{i=1}^{n} \sum_{l=0}^{\infty} \sum_{k=0}^{\infty} \sum_{d=0}^{\infty} (1 - \delta_{i}) \frac{a^{2}b}{4\alpha} (-1)^{l+k+d} \binom{2a-2}{l} \binom{b-2}{d} \binom{ad}{k} \frac{2^{-(l+k)}}{-(l+k+d+1)^{2}}\right]$$

$$\begin{split} E\Big[u_{\mu\mu}\Big] &= -[\sum_{i=1}^n \delta_i [\sum_{j=0}^\infty \frac{a^2b}{\alpha^2} (-1)^j \binom{b-3}{j} \frac{(-2)^{-a(j+3)}}{(j+3)} - \sum_{l=0}^\infty \frac{a^2b}{\alpha^3} (-1)^l \binom{b-3}{l} \frac{(-2)^{-a(l+2)}}{(l+2)}] + \\ &\sum_{i=1}^n (1-\delta_i) [\frac{-ab(a-1)}{8\alpha^2} \sum_{d=0}^\infty \sum_{k=0}^\infty \sum_{f=0}^\infty (-1)^{f+k+d} \binom{a-3}{d} \binom{b-1}{k} \binom{ak}{f} \frac{2^{-(d+f)}}{f+d+3} - \frac{a^3b(b-1)}{8\alpha^2} \sum_{u=0}^\infty \sum_{w=0}^\infty \sum_{x=0}^\infty (-1)^{u+w+x} \binom{a-2}{u} \binom{b-1}{w} \binom{aw}{x} \frac{2^{-(u+x)}}{u+x+2} - \frac{a^3b(b-1)}{8\alpha^2} \sum_{m=0}^\infty \sum_{n=0}^\infty \sum_{n=0}^\infty (-1)^{m+n+y} \binom{3a-3}{m} \binom{b-3}{n} \binom{an}{y} \frac{2^{-(n+m+y)}}{m+y+3} - \frac{ba^2(a-1)(b-1)}{8\alpha^2} \sum_{c=0}^\infty \sum_{g=0}^\infty \sum_{e=0}^\infty (-1)^{e+g+c} \frac{2^{-(c+g)}}{c+g+3} + \frac{a^2b(b-1)}{4\alpha^2} \sum_{h=0}^\infty \sum_{o=0}^\infty \sum_{p=0}^\infty (-1)^{0+p+h} \binom{2a-2}{h} \binom{b-2}{o} \binom{ao}{p} \frac{2^{-(p+h)}}{p+h+2}] \end{split}$$

$$\begin{split} E[u_{aa}] &= -\left[\sum_{l=1}^{n} \delta_{l} [-\sum_{j=0}^{\infty} (-1)^{j} \binom{b-1}{j} \frac{ab}{a} \frac{2^{-(a+j)}}{(a+j)^{2}} - \frac{2^{-3a}a^{2}b(b-1)}{a} \sum_{k=0}^{\infty} \sum_{m=1}^{\infty} \sum_{l=0}^{\infty} \sum_{f=0}^{\infty} \sum_{n=1}^{\infty} \frac{(-1)^{k+l+f}}{nm} \binom{b-3}{k} \binom{n}{l} \binom{m}{f} \frac{1}{a(k+3)+f+l} - \frac{ab(b-1)}{a} \sum_{s=0}^{\infty} \sum_{t=1}^{\infty} \sum_{u=0}^{\infty} \frac{(-1)^{u+s}}{t} \binom{b-2}{s} \binom{t}{u} \frac{2^{-a(s+2)}}{a(s+u+2)} + \frac{a^{2}b(b-1)}{a} \sum_{v=0}^{\infty} \sum_{w=1}^{\infty} \sum_{y=0}^{\infty} \sum_{0}^{\infty} \sum_{w=1}^{\infty} \sum_{y=0}^{\infty} \frac{(-1)^{y+v}}{w} \binom{b-2}{v} \binom{w}{y} \frac{2^{-a(v+2)}}{a(v+y+2)} - \frac{2^{-2a}a^{2}b(b-1)}{a} \sum_{x=1}^{\infty} \sum_{z=0}^{\infty} \sum_{0}^{\infty} \sum_{d=0}^{\infty} \frac{(-1)^{d+z}}{x} \binom{b-2}{d} \binom{x}{z} \frac{2^{-a(j+2)}}{a(v+y+2)} \right] + \\ \sum_{l=1}^{n} (1-\delta_{l}) [-\frac{ab}{a} \sum_{e=0}^{\infty} \sum_{c=0}^{\infty} \sum_{0}^{\infty} \sum_{0}^{\infty} \sum_{h=1}^{\infty} \sum_{0=0}^{\infty} \frac{(-1)^{e+c+g}}{a(v+b+c)} \binom{a-2}{v} \binom{b-2}{c} \binom{ac}{c} \binom{b-1}{c} \binom{ac}{d} \binom{o-1}{h} \frac{2^{-(e+g+2)}}{e+g+h+2} + \frac{ab(b-1)}{a} \sum_{x'=0}^{\infty} \sum_{y'} \sum_{0'=0}^{\infty} \sum_{c'=0}^{\infty} \sum_{c'=0}^{\infty} \sum_{c'=0}^{\infty} \sum_{0}^{\infty} \frac{(-1)^{x'+c'+c'+c'}}{y'} \binom{a-2}{c} \binom{y'}{c} \binom{ac'}{c'} \binom{b-2}{c'} \left[\frac{2^{-(x'+c''+2)}}{x'+c''+o'+1} \right] - \frac{a^{2}b(b-1)}{a} \sum_{d'=0}^{\infty} \sum_{0}^{\infty} \sum_{$$

$$\begin{split} & \sum_{i=0}^{\infty} \sum_{p'=1}^{\infty} \sum_{t=0}^{\infty} \frac{a^3 b(b-1)}{a} \frac{(-1)^{i+p}}{p'} \binom{b}{p} \binom{b}{t} \binom{p'}{a} \frac{2^{-a(p+2)}}{a(p+2)+t} \end{vmatrix} + \\ & \sum_{i=0}^{\infty} (1-\delta_l) [-\frac{ab}{a^2} \sum_{p'=0}^{\infty} \sum_{j=0}^{\infty} \sum_{j$$

$$\frac{a^2b}{a^2} \sum_{k=0}^{\infty} \sum_{n=1}^{\infty} \sum_{l=0}^{(-1)^{l+k}} \binom{b-1}{k} \binom{n}{l} \frac{2^{-(ak+a-1)}}{a(k+1)+l} + \\ \frac{a^2b(b-1)}{a^2} \sum_{p=0}^{\infty} \sum_{s=0}^{\infty} \sum_{m=1}^{\infty} \frac{(-1)^{p+k}}{m} \binom{p}{n} \binom{b-2}{s} \frac{2^{-a(s+2)+1}}{a(s+2)+p} + \\ \frac{a^3b(b-1)}{a^2} \sum_{p=0}^{\infty} \sum_{r=0}^{\infty} \sum_{r=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{r+p'}}{s} \binom{b-3}{s} \binom{o}{p'} \frac{2^{-ax-3a+1}}{3a+ax+p'} + \\ \frac{a^3b(b-1)}{a} \sum_{y=1}^{\infty} \sum_{y'=0}^{\infty} \sum_{z'=0}^{\infty} \frac{(-1)^{x'+y'}}{y} \binom{y'}{y'} \binom{b-2}{s^2} \frac{2^{-ax'-2a+1}}{y'+a(x'+2)} + \\ \sum_{l=1}^{n} (1-\delta_l) \left[\frac{ab}{a^2} \sum_{j'=0}^{\infty} \sum_{c'=0}^{\infty} \sum_{c'=0}^{\infty} \sum_{c''=0}^{\infty} (-1)^{j'+c'+c''} \binom{a-1}{j'} \binom{b-1}{c'} \binom{ac'}{c''} \frac{2^{-(j'+c''+1)}}{(j'+c''+1)} - \\ \frac{ab}{a^2} \sum_{l'=0}^{\infty} \sum_{q=0}^{\infty} \sum_{q=0}^{\infty} \sum_{c''=0}^{\infty} \sum_{c''=0}^{\infty} \sum_{c''=0}^{\infty} (-1)^{j'+c'+c''} \binom{a-1}{j'} \binom{b-1}{c'} \binom{ac'}{a''} \binom{d-1}{a'} \binom{ac'}{l'+e+q''+q'+1} + \\ \frac{ab(a-1)}{a^2} \sum_{s=1}^{\infty} \sum_{y''=0}^{\infty} \sum_{e=0}^{\infty} \sum_{e'=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{s''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \binom{a-1}{j''+e+e'+e''} \binom{s}{s'} \binom{a-2}{e-2} \binom{b-1}{c'} \binom{ae'}{a'} \frac{2^{-(e+e''+1)}}{(e+e''+y''+2)} - \\ \frac{a^2b(b-1)}{a^2} \sum_{y''=1}^{\infty} \sum_{j''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \binom{a-1}{j''+e+e'+e''} \binom{s}{s'} \binom{a-2}{s'} \binom{b-1}{e'} \binom{ae'}{s'} \frac{2^{-(e+e''+1)}}{(e+e''+y''+2)} + \\ \frac{ab(a-1)}{a^2} \sum_{y''=1}^{\infty} \sum_{j''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e'''=0}^{\infty} \binom{a-1}{j''+e+e'+e'+e''} \binom{s}{s'} \binom{s}{s'} \binom{a-2}{s'} \binom{b-1}{s'} \binom{ah'}{s'} \frac{2^{-(e+e''+1)}}{(e+e''+y''+2)} + \\ \frac{ab(a-1)}{a^2} \sum_{y''=1}^{\infty} \sum_{s''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e'''=0}^{\infty} \sum_{e'''=0}^{\infty} \binom{a-1}{s''} \binom{s''+e+e'+e''+e''}{s''} \binom{s'}{s'} \binom{a-2}{s'} \binom{b-1}{s'} \binom{ae'}{s''} \frac{2^{-(e+e''+1)}}{(e+e''+y''+2)} + \\ \frac{ab(a-1)}{a^2} \sum_{s''=1}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e''=0}^{\infty} \sum_{e'''=0}^{\infty} \binom{a-1}{s''} \binom{s''+e+e'+e'+e''}{s''} \binom{s'}{s'} \binom{a-2}{s'} \binom{b-1}{s''} \binom{ah'}{s''} \binom{a-2}{s''} \binom{ah'}{s''} \binom{a-2}{s''} \binom{ah'}{s''} \binom{a-2}{s''} \binom{ah'}{s''} \binom{ah'}{s''} \binom{ah'}{s''} \binom{ah'}{s''} \binom{ah'}{s''} \binom{ah'}{s''} \binom{ah'}{s''} \binom{ah'$$

1. Renyi and Shannon entropies

The entropy of a random variable is a measure of variation of the uncertainty. In information theory; the Renyi entropy generalizes the Shannon entropy. The entropy is important in ecology, statistics, engineering, and in quantum information. The Renyi entropy of order γ , where $\gamma \ge 0$ and $\gamma \ne 1$, is defined as

$$H_{\gamma} = \frac{1}{1-\gamma} Log[A(\gamma)]$$
, where $A[\gamma] = \int f^{\gamma}(t)dt$, so from (6),

$$f^{\gamma}(t) = \begin{cases} \left(\frac{2^{-a}ab(e^{\frac{-\mu+t}{\alpha}})^{a}(1-2^{-a}(e^{\frac{-\mu+t}{\alpha}})^{a})^{-1+b}}{\alpha}\right)^{\gamma} & t < \mu \\ \left(\frac{abe^{-\frac{-\mu+t}{\alpha}}(1-\frac{1}{2}e^{-\frac{-\mu+t}{\alpha}})^{-1+a}(1-(1-\frac{1}{2}e^{-\frac{-\mu+t}{\alpha}})^{a})^{-1+b}}{2\alpha}\right)^{\gamma} & t \ge \mu \end{cases}$$

The Kumaraswamy- Laplace Distribution

$$\begin{split} \mathbf{A}[\gamma] = & \int_{-\infty}^{0} \left(\frac{2^{-a} a b (e^{\frac{-\mu + t}{\alpha}})^{a} (1 - 2^{-a} (e^{\frac{-\mu + t}{\alpha}})^{a})^{-1 + b}}{\alpha} \right)^{\gamma} dt + \\ & \int_{0}^{\infty} \left(\frac{a b e^{-\frac{-\mu + t}{\alpha}} (1 - \frac{1}{2} e^{-\frac{-\mu + t}{\alpha}})^{-1 + a} (1 - (1 - \frac{1}{2} e^{-\frac{-\mu + t}{\alpha}})^{a})^{-1 + b}}{2\alpha} \right)^{\gamma} dt \end{split}$$

$$\begin{split} \mathbf{A}[\gamma] &= & \frac{(2^{-a}ab)^{\gamma}}{\alpha^{1-\gamma}} \sum_{j=0}^{\infty} \frac{(-1)^{j}}{2^{aj}} \binom{\gamma(b-1)-1}{j} \frac{1}{(a\gamma+j)} + \\ & \sum_{l=0}^{\infty} \sum_{m=0}^{\infty} \sum_{k=0}^{\infty} \frac{(ab)^{\gamma}}{2^{k+m+\gamma}} \left(\frac{1}{\alpha}\right)^{1-\gamma} \frac{(-1)^{k+l+m}}{k+m+\gamma} \binom{\gamma(a-1)-1}{k} \binom{\gamma(b-1)-1}{l} \binom{al-1}{m} \end{split}$$

Hence, the Renyi entropy reduces to

$$\begin{split} &H_{\gamma}(t) \\ &= \frac{1}{1-\gamma} log[\frac{(2^{-a}ab)^{\gamma}}{\alpha^{1-\gamma}} \sum\nolimits_{j=0}^{\infty} \frac{(-1)^{j}}{2^{aj}} \binom{\gamma(b-1)-1}{j} \frac{1}{(a\gamma+j)} \\ &+ \sum\nolimits_{l=0}^{\infty} \sum\nolimits_{m=0}^{\infty} \sum\nolimits_{k=0}^{\infty} \frac{(ab)^{\gamma}}{2^{k+m+\gamma}} \binom{1}{\alpha}^{1-\gamma} \frac{(-1)^{k+l+m}}{k+m+\gamma} \binom{\gamma(a-1)-1}{k} \binom{\gamma(b-1)-1}{l} \binom{al-1}{m}] \end{split}$$

The Shannon entropy is defined by

$$H(t) = E[-Log[f(t)]] = -[\int_{-\infty}^{\mu} \log[f(t)]f(t)dt + \int_{\mu}^{\infty} \log[f(t)]f(t)dt]$$

From equation (6), the Shannon entropy can be written as

$$-b \log \left[\frac{2^{-a}ab}{\alpha}\right] \sum_{j=0}^{\infty} 2^{-a(j+1)} {b-1 \choose j} \frac{(-1)^j}{(j+1)}$$

$$+ ab \sum_{l=1}^{\infty} \sum_{k=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{k+m}}{l} {k \choose k} {b-1 \choose m} \frac{2^{-a}(m+1)}{(m+1)+k}$$

$$+ (b-1) \sum_{n=1}^{\infty} \sum_{r=0}^{\infty} (-1)^r {b-1 \choose r} \frac{2^{-a(n+r+1)}}{n} * \frac{b}{(n+1)+r}$$

$$- \frac{ab}{2} \log \left[\frac{ab}{2a}\right] \sum_{x=0}^{\infty} \sum_{y=0}^{\infty} \sum_{i=0}^{\infty} (-1)^{i+y+x} {a-1 \choose x} {b-1 \choose y} {ay \choose i} \frac{2^{-(i+x)}}{i+x+1}$$

$$+ \frac{ab}{2} \sum_{c=0}^{\infty} \sum_{d=0}^{\infty} \sum_{p=0}^{\infty} \sum_{f=0}^{\infty} \sum_{g=0}^{\infty} \frac{(-1)^{d+e+f+g}}{c} {c \choose d} {a-1 \choose f} {af \choose g} \frac{2^{-(e+g)}}{e+g+d+1}$$

$$+ \sum_{n=1}^{\infty} \sum_{h=0}^{\infty} \sum_{p=0}^{\infty} \sum_{q=0}^{\infty} ab(a)$$

$$- 1)(-1)^{h+p+q} {a-1 \choose h} {b-1 \choose p} {ap \choose q} \frac{2^{-(h+q)}}{n^c+h+q+1}$$

$$+ \sum_{m=1}^{\infty} \sum_{s=0}^{\infty} \sum_{t=0}^{\infty} \sum_{r=0}^{\infty} \sum_{t=0}^{\infty} (-1)^{s+t+v} {a(m^c+1) \choose s} {b-1 \choose t} {at \choose v} \frac{2^{-(v+s)}}{s+v+1}$$

5. Application

In this section we fit the Kumaraswamy-Laplace (KL), Beta-Laplace (BL)[see Cordeiro and Lemonte (2011)], and Laplace distribution models to the data in table[1], and compare it with our models. First, we estimate unknown parameters of the models by the maximum likelihood method and then we obtain the values of Akaike Information criterion (AIC) and Bayesian Information Criterion (BIC). A summary of computations is given in table [2]

Table 1:

-5.4	-1.5	-8.5	-6.1	-4.7	-0.5	-4.1	-2.8	-1.7	-4.09
-8.2	-9.6	-3.4	-0.6	-6.2	-9.8	-2.4	-9.6	-5.5	-7.5
-1.6	-4.6	-4.2	-8.1	-6.8	-2.04	-5.8	-0.8	-3.1	-4.4
-2.2	-8.5	-4.07	-6.9	-8.1	-9.8	-9.4	-8.4	-1.7	-3.7
0.9	5.2	0.3	1.2	2.1	4.3	3.3	2.7	0.4	5.3
3.8	6.5	1.8	8.2	1.1	1.06	2.06	1.08	2.9	6.6
5.5	5.3	4.9	0.7	6.7	8.4	4.4	4.2	4.6	3.9
5.1	0.6	2.4	7.1	3.3	0.9	4.02	9.5	3.1	1.5

Table 2: MLE of the model parameters, Log-Likelihood functions $[l(\widehat{\theta})]$ and AIC and BIC statistics

Distribution	Estimates			$l(\hat{ heta})$	AIC	BIC	
	a	b	α	μ			
KL	0.653167	0.578913	1.8755	0.5804	-268.129	544.25	553.78
BL	2.149×10^{8}	2.149×10^{8}	1.156×10^9	0	-947.848	1901.696	1908.8420

As we see from the results presented in Table 2, with $\hat{\theta} = (\hat{a}, \hat{b}, \hat{\alpha}, \hat{\mu})$, the KL model with the minimum values of AIC and BIC gives a better fit than BL model, so the new KL model provides consistently better fit than the other model.

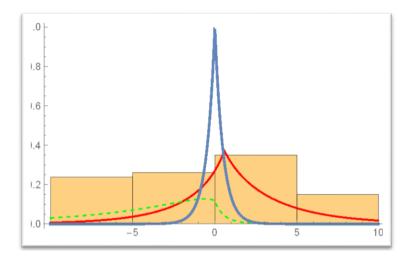


Fig. 2. Estimate pdf of the KL, BL and the Laplace distribution for the real data, where---BL, — KL, — Laplace Distributions.

References

- 1. Akinsete, F. Famoye, C. Lee, (2008). The beta-Pareto distribution, Statistics 42, 547–563.
- 2. A. Akinsete, F. Famoye, and C. Lee (2014). The Kumaraswamy geometric distribution, Journal of statistical distributions and applications, 2014, 1-17
- 3. H. Bidramand V. Nekoukhou (2013). Double bounded Kumaraswamy power series class of distributions, SORT 37 (2), 211-230.
- 4. G. M. Cordeiro and A.J. Lemonte (2011). The beta Laplace distribution, Statistics and Probability Letters 81(8), 973-982.
- 5. R. D. Gupta and D. Kundu, (2001). Exponentiated exponential family: an alternative to gamma and Weibull. Biometrical Journal 43, 117-130.
- 6. M. C. Jones (2009). Kumaraswamy distribution: A beta type Distribution with some tractability advantages, Statistical Methodology, 6(1), 70-81.
- 7. L. Kong, C. Lee and J.H. Sepanski (2007). On the properties of beta Gamma distributions, Journal of Modern Applied Statistical Methods 6, 173–187.
- 8. P. Kumaraswamy (1980). Generalized probability density-function for double-bounded random-processes. Journal of Hydrology, 462, 79–88.
- 9. Mahmoud, M., Nassar, M. M. and Aefa, M. A. (2015). "A new generalization of Kumaraswamy Lindley distribution" Journal of Statistics- Advances in Theory and Applications, 14, 69-105.
- 10. G. S. Mudholkar, and D. K. Srivastava, (1993). Exponentiated Weibull family for analyzing bathtub failure-rate data. IEEE Transactions on Reliability 42, 299-302.
- 11. S. Nadarajah and S. Eljabri, (2013). The Kumaraswamy GP Distribution, Journal of Data Science 11, 739-766.
- 12. S. Nadarajah and S. Kotz, (2004). The beta Gumbel distribution, Mathematics and Probability for Engineers 10, 323–332.
- 13. S. Nadarajah and S. Kotz, (2005). The beta exponential distribution, Reliability Engineering and System Safety 91, 689–697.
- 14. M. M. Nassarand F. H. Eissa (2003). "On the exponentiated Weibull distribution" Communications in Statistics Theory and Methods, Vol. 32, No. 7, 1317-1336.
- 15. M. M. Nassarand F. H. Eissa (2004) "Bayesian estimation for the exponentiated Weibull model" Communications in Statistics Theory and Methods, Vol. 33, No. 10, 2343-2362.
- 16. M.M. Nassar, N.K. Nada, (2011). The beta generalized Pareto distribution, Journal of Statistics Advances in Theory and Applications 6, 1–17.
- 17. M. M. Nassarand N. K. Nada (2012). "A new generalization of the exponential geometric distribution" Journal of Statistics- Advances in Theory and Applications, 7, 25-48.
- 18. M. M. Nassarand A. Elmasry (2012). "A study of generalized Logistic distributions" Journal of Egyptian Mathematical Society, 20(2), 126-133.

- 19. M. M. Nassar and N. K. Nada (2013). "A new generalization of the Pareto geometric distribution" Journal of Egyptian Mathematical Society, 21(2), 148-155.
- 20. R.R. Pescim, C.G.B. Demetrio, G.M. Cordeiro, E.M. M. Ortega and M. R. Urbano (2010). The beta generalized half-normal distribution, Computational Statistics and Data Analysis 54, 945–957.
- 21. T. V. de Santana, E.M. M. Ortega, G. M. Cordeiro (2012). The Kumaraswamy-LOG-Logistic Distribution, Journal of Statistics Theory and Applications, 11 (3), 265-291.
- 22. W.B. Souza, A.H.S. Santos, G.M. Cordeiro, (2010). The beta generalized exponential distribution, Journal of Statistical Computation and Simulation 80, 159–172.
- 23. L. Xiaohu, H. Yanyan and Z. Xueyan (2011). The Kumaraswamy Binomial Distribution, Chinese Journal of Applied Probability and Statistics, 27 (5), 511-521.