

Some Improved Modified Ratio Estimators Based on Decile Mean of an Auxiliary Variable

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Abstract

The estimators developed so far regarding the study under consideration use the conventional measures of central tendencies, i.e. the mean, the median, the quartile mean etc., and comment on their properties. However, Sohail et al. (2012) have proposed decile mean as a measure of central tendency and have proved that it outperforms the conventional measures of central tendency. In this study, we have attempted to use the decile mean instead of the conventional measures suggested in previous studies. Also, we have used decile mean, population correlation coefficient, coefficient of variation and the linear combinations of auxiliary variable and investigated the properties associated with the proposed estimator. Theoretically, mean square error equations of all proposed ratio estimators are obtained and the efficiency conditions are derived. This study has been verified numerically.

Keywords: Auxiliary variable; Coefficient of variation; Decile mean; Mean squared error; Population correlation coefficient.

1. Introduction

Now a day, the information obtained through auxiliary variables was largely discussed in survey sampling. The auxiliary variables are generally connected with the study variables and we may take up this information in different forms such as ratio, product and regression etc. The auxiliary information may be accessible from several sources such as similar studies in past, economic reports, national census etc.

The ratio and regression estimators are used to improve the efficiency of the simple random sampling without replacement (SRSWOR) sample mean when there is a positive correlation exist between study variable and an auxiliary variable under certain conditions (see for example Cochran (1977) and Murthy (1967)).

The problem of constructing efficient estimators for the population mean has been widely discussed by Several researcher such as Cochran (1940), Kadilar and Cingi (2004, 2006)), Murthy (1967), Rao (1991), Singh and Tailor (2003), Sisodia and Dwivedi (1981), Subramani and Kumarapandiyan (2012a, 2012b, 2012c, 2012d), Upadhyaya and Singh (1999), Yan and Tian (2010), Abid et al. (2016a, 2016b, 2016c) are developed efficient estimators for the population mean based on the auxiliary obtained in the form of ratio estimator.

Let $Z = \{Z_1, Z_2, Z_3, \dots, Z_N\}$ be the N different and particular units from a finite population and Y be the study variable with value Y_i obtained from Z_i , $i = 1, 2, \dots, N$. The purpose is to estimation population mean $\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i$ on the base of a random sample.

The notations considered in this article are as follows:

NOMENCLATURE

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N	Population size
n	Sample size
$f = n/N$	Sampling fraction
Y	Study variable
X	Auxiliary variable
\bar{X}, \bar{Y}	Population means
\bar{x}, \bar{y}	Sample means
x, y	Sample totals
S_x, S_y	Population standard deviations
S_{xy}	Population covariance between X and Y
C_x, C_y	Coefficient of variation
$B(.)$	Bias of the Estimator
$MSE(.)$	Mean square error of the estimator
$\hat{\bar{Y}}_i$	Existing modified ratio estimator of \bar{Y}
$\hat{\bar{Y}}_{pj}$	Proposed modified ratio estimator of \bar{Y}
M_d	Median of auxiliary variable
$QD = \frac{Q_3 - Q_1}{2}$	Quartile Deviation
$G = \frac{4}{N-1} \sum_{i=1}^N \left(\frac{2i-N-1}{2N} \right) X_{(i)}$	Gini's Mean Difference
$D = \frac{2\sqrt{\pi}}{N(N-1)} \sum_{i=1}^N \left(i - \frac{N+1}{2} \right) X_{(i)}$	Downton's method
$S_{pw} = \frac{\sqrt{\pi}}{N^2} \sum_{i=1}^N (2i - N - 1) X_{(i)}$	Probability Weighted Moments
$DM = \frac{D_1 + D_2 + \dots + D_9}{9}$	Decile Mean

Subscript

- i For existing estimators
 j For proposed estimators

Greek

ρ Coefficient of correlation

$\beta_1 = \frac{N \sum_{i=1}^N (X_i - \bar{X})^3}{(N-1)(N-2)S^3}$ Coefficient of skewness of auxiliary variable

$\beta_2 = \frac{N(N+1) \sum_{i=1}^N (X_i - \bar{X})^4}{(N-1)(N-2)(N-3)S^4} - \frac{3(N-1)^2}{(N-2)(N-3)}$ Coefficient of kurtosis of auxiliary variable

$b = \frac{S_{xy}}{S_x^2}$ Regression coefficient of Y on X

The ratio estimator based on the population mean, \bar{Y} , is defined as

$$\hat{Y}_r = \frac{\bar{y}}{\bar{x}} \bar{X} \quad (1)$$

The bias, related constant and the mean squared error (MSE) of the ratio estimator are respectively given by

$$B(\hat{Y}_r) = \frac{(1-f)}{n} \frac{1}{\bar{X}} (RS_x^2 - \rho S_x S_y), \quad R = \frac{\bar{Y}}{\bar{X}}, \quad MSE(\hat{Y}_r) = \frac{(1-f)}{n} (S_y^2 + R^2 S_x^2 - 2R\rho S_x S_y).$$

The rest of the study is presented as follows: Section 2 gives a explanation of the existing estimators. The construction and the efficiency comparison of the suggested estimator with the existing estimators are obtainable in Section 3. Section 4 contains the numerical comparison of the proposed and existing estimators. Finally, conclusions of the study are given in Section 5.

2. The existing modified ratio estimators

Kadilar and Cingi (2004, 2006) has been proposed some linear regression type ratio estimators as follows:

$$\begin{aligned} \hat{Y}_1 &= \frac{\bar{y}+b(\bar{X}-\bar{x})}{\bar{x}} \bar{X}, \quad \hat{Y}_2 = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+C_x)} (\bar{X} + C_x), \quad \hat{Y}_3 = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+\beta_2)} (\bar{X} + \beta_2), \\ \hat{Y}_4 &= \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\beta_2+C_x)} (\bar{X}\beta_2 + C_x), \quad \hat{Y}_5 = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+\beta_2)} (\bar{X}C_x + \beta_2), \quad \hat{Y}_6 = \frac{\bar{y}+b(\bar{X}-\bar{x})}{\bar{x}} (\bar{X} + \rho), \\ \hat{Y}_7 &= \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+\rho)} (\bar{X}C_x + \rho), \\ \hat{Y}_8 &= \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\rho+C_x)} (\bar{X}\rho + C_x), \quad \hat{Y}_9 = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\beta_2+\rho)} (\bar{X}\beta_2 + \rho), \quad \hat{Y}_{10} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\rho+\beta_2)} (\bar{X}\rho + \beta_2). \end{aligned}$$

Yan and Tian (2010) suggested some modified linear regression type ratio estimators as follows:

$$\hat{Y}_{11} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+\beta_1)} (\bar{X} + \beta_1), \quad \hat{Y}_{12} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\beta_1+\beta_2)} (\bar{X}\beta_1 + \beta_2).$$

The Subramani and Kumarapandiyan (2012a, 2012b, 2012c) suggested estimators based on population median, skewness, kurtosis and coefficient of variation of an auxiliary variable are shown below:

$$\hat{Y}_{13} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+M_d)}(\bar{X}+M_d), \quad \hat{Y}_{14} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+M_d)}(\bar{X}C_x+M_d), \quad \hat{Y}_{15} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\beta_1+M_d)}(\bar{X}\beta_1+M_d),$$

$$\hat{Y}_{16} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\beta_2+M_d)}(\bar{X}\beta_2+M_d).$$

Jeelani et al. (2013) developed an estimator based on quartile deviation and skewness. Jeelani et al. (2013) can be written as,

$$\hat{Y}_{17} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\beta_1+QD)}(\bar{X}\beta_1+QD).$$

Abid et al. (2016c) developed some modified ratio estimators by using non-conventional measures of dispersions. Abid et al. (2016c) estimators can be written as,

$$\hat{Y}_{18} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+G)}(\bar{X}+G), \quad \hat{Y}_{19} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\rho+G)}(\bar{X}\rho+G), \quad \hat{Y}_{20} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+G)}(\bar{X}C_x+G),$$

$$\hat{Y}_{21} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+D)}(\bar{X}+D), \quad \hat{Y}_{22} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\rho+D)}(\bar{X}\rho+D), \quad \hat{Y}_{23} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+D)}(\bar{X}C_x+D),$$

$$\hat{Y}_{24} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+S_{pw})}(\bar{X}+S_{pw}), \quad \hat{Y}_{25} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\rho+S_{pw})}(\bar{X}\rho+S_{pw}), \quad \hat{Y}_{26} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+S_{pw})}(\bar{X}C_x+S_{pw}).$$

The constants, biases and MSEs of the estimators developed by Kadilar and Cingi (2004, 2006), Yan and Tian (2010), Subramani and Kumarapandiyan (2012a, 2012b, 2012c), Jeelani et al. (2013) and Abid et al. (2016c) are given below, respectively,

$$R_1 = \frac{\bar{Y}}{\bar{X}}, \quad R_2 = \frac{\bar{Y}}{(\bar{X}+C_x)}, \quad R_3 = \frac{\bar{Y}}{(\bar{X}+\beta_2)}, \quad R_4 = \frac{\bar{Y}\beta_2}{(\bar{X}\beta_2+C_x)}, \quad R_5 = \frac{\bar{Y}C_x}{(\bar{X}C_x+\beta_2)}, \quad R_6 = \frac{\bar{Y}}{\bar{X}+\rho},$$

$$R_7 = \frac{\bar{Y}C_x}{(\bar{X}C_x+\rho)}, \quad R_8 = \frac{\bar{Y}\rho}{(\bar{X}\rho+C_x)}, \quad R_9 = \frac{\bar{Y}\beta_2}{(\bar{X}\beta_2+\rho)}, \quad R_{10} = \frac{\bar{Y}\rho}{(\bar{X}\rho+\beta_2)}, \quad R_{11} = \frac{\bar{Y}}{(\bar{X}+\beta_1)},$$

$$R_{12} = \frac{\bar{Y}\beta_1}{(\bar{X}\beta_1+\beta_2)}, \quad R_{13} = \frac{\bar{Y}}{(\bar{X}+M_d)}, \quad R_{14} = \frac{\bar{Y}C_x}{(\bar{X}C_x+M_d)}, \quad R_{15} = \frac{\bar{Y}\beta_1}{(\bar{X}\beta_1+M_d)}, \quad R_{16} =$$

$$\frac{\bar{Y}\beta_2}{(\bar{X}\beta_2+M_d)}, \quad R_{17} = \frac{\bar{Y}\beta_1}{(\bar{X}\beta_1+QD)}, \quad R_{18} = \frac{\bar{Y}}{(\bar{X}+G)},$$

$$R_{19} = \frac{\bar{Y}\rho}{(\bar{X}\rho+G)}, \quad R_{20} = \frac{\bar{Y}C_x}{(\bar{X}C_x+G)}, \quad R_{21} = \frac{\bar{Y}}{(\bar{X}+D)}, \quad R_{22} = \frac{\bar{Y}\rho}{(\bar{X}\rho+D)}, \quad R_{23} =$$

$$\frac{\bar{Y}C_x}{(\bar{X}C_x+D)}, \quad R_{24} = \frac{\bar{Y}}{(\bar{X}+S_{pw})}, \quad R_{25} = \frac{\bar{Y}\rho}{(\bar{X}\rho+S_{pw})}, \quad R_{26} = \frac{\bar{Y}C_x}{(\bar{X}C_x+S_{pw})}.$$

$$B(\hat{Y}_1) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_1^2, \quad B(\hat{Y}_2) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_2^2, \quad B(\hat{Y}_3) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_3^2, \quad B(\hat{Y}_4) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_4^2,$$

$$B(\hat{Y}_5) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_5^2, \quad B(\hat{Y}_6) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_6^2, \quad B(\hat{Y}_7) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_7^2, \quad B(\hat{Y}_8) =$$

$$\frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_8^2, \quad B(\hat{Y}_9) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_9^2, \quad B(\hat{Y}_{10}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_{10}^2, \quad B(\hat{Y}_{11}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_{11}^2,$$

$$B(\hat{Y}_{12}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_{12}^2,$$

$$B(\hat{Y}_{13}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_{13}^2, \quad B(\hat{Y}_{14}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_{14}^2, \quad B(\hat{Y}_{15}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}} R_{15}^2, \quad B(\hat{Y}_{16}) = \frac{(1-f)}{n} \frac{S_x^2}{\bar{Y}},$$

$$\begin{aligned}
 B(\hat{Y}_{17}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{17}^2, & B(\hat{Y}_{18}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{18}^2, & B(\hat{Y}_{19}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{19}^2, & B(\hat{Y}_{20}) &= \\
 & \frac{(1-f)S_x^2}{n\bar{Y}}R_{20}^2, & B(\hat{Y}_{21}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{21}^2, & B(\hat{Y}_{22}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{22}^2, & B(\hat{Y}_{23}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{23}^2, \\
 B(\hat{Y}_{24}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{24}^2, & B(\hat{Y}_{25}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{25}^2, & B(\hat{Y}_{26}) &= \frac{(1-f)S_x^2}{n\bar{Y}}R_{26}^2. \\
 \\
 MSE(\hat{Y}_1) &= \frac{(1-f)}{n}(R_1^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_2) &= \frac{(1-f)}{n}(R_2^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_3) &= \frac{(1-f)}{n}(R_3^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_4) &= \frac{(1-f)}{n}(R_4^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_5) &= \frac{(1-f)}{n}(R_5^2S_x^2 + S_y^2(1-\rho^2)), & E(\hat{Y}_6) &= \frac{(1-f)}{n}(R_6^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_7) &= \frac{(1-f)}{n}(R_7^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_8) &= \frac{(1-f)}{n}(R_8^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_9) &= \frac{(1-f)}{n}(R_9^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{10}) &= \frac{(1-f)}{n}(R_{10}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{11}) &= \frac{(1-f)}{n}(R_{11}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{12}) &= \frac{(1-f)}{n}(R_{12}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{13}) &= \frac{(1-f)}{n}(R_{13}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{14}) &= \frac{(1-f)}{n}(R_{14}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{15}) &= \frac{(1-f)}{n}(R_{15}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{16}) &= \frac{(1-f)}{n}(R_{16}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{17}) &= \frac{(1-f)}{n}(R_{17}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{18}) &= \frac{(1-f)}{n}(R_{18}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{19}) &= \frac{(1-f)}{n}(R_{19}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{20}) &= \frac{(1-f)}{n}(R_{20}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{21}) &= \frac{(1-f)}{n}(R_{21}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{22}) &= \frac{(1-f)}{n}(R_{22}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{23}) &= \frac{(1-f)}{n}(R_{23}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{24}) &= \frac{(1-f)}{n}(R_{24}^2S_x^2 + S_y^2(1-\rho^2)), \\
 MSE(\hat{Y}_{25}) &= \frac{(1-f)}{n}(R_{25}^2S_x^2 + S_y^2(1-\rho^2)), & MSE(\hat{Y}_{26}) &= \frac{(1-f)}{n}(R_{26}^2S_x^2 + S_y^2(1-\rho^2)).
 \end{aligned}$$

3. The Suggested modified ratio estimators

In this section, we have suggested some modified ratio type estimators using the auxiliary information on population decile mean, population coefficient of variation and population correlation coefficient for the estimation of population mean of a variable of interest. Sohail et al. (2012) showed that the decile mean perform better than the conventional measures of locations such as mean, median, mode in the presence of extreme values. So, it is better to use decile mean instead of mean, median and mode when the extreme values are present in the data sets. The proposed estimators are shown below,

$$\hat{Y}_{p1} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}+DM)}(\bar{X} + DM), \quad \hat{Y}_{p2} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}C_x+G)}(\bar{X}C_x + DM), \quad \hat{Y}_{p3} = \frac{\bar{y}+b(\bar{X}-\bar{x})}{(\bar{x}\rho+G)}(\bar{X}\rho + DM).$$

whereas the bias, constant and MSE of these proposed estimators are respectively as under:

$$B(\hat{Y}_{p1}) = \frac{(1-f)S_x^2}{n\bar{Y}}R_{p1}^2, \quad B(\hat{Y}_{p2}) = \frac{(1-f)S_x^2}{n\bar{Y}}R_{p2}^2, \quad B(\hat{Y}_{p3}) = \frac{(1-f)S_x^2}{n\bar{Y}}R_{p3}^2.$$

$$R_{p1} = \frac{\bar{Y}}{(\bar{X}+G)}, \quad R_{p2} = \frac{\bar{Y}C_x}{(\bar{X}C_x+DM)}, \quad R_{p3} = \frac{\bar{Y}\rho}{(\bar{X}\rho+DM)}.$$

$$MSE(\hat{\bar{Y}}_{p1}) = \frac{(1-f)}{n} \left(R_{p1}^2 S_x^2 + S_y^2 (1 - \rho^2) \right), \quad MSE(\hat{\bar{Y}}_{p2}) = \frac{(1-f)}{n} \left(R_{p2}^2 S_x^2 + S_y^2 (1 - \rho^2) \right),$$

$$MSE(\hat{\bar{Y}}_{p3}) = \frac{(1-f)}{n} \left(R_{p3}^2 S_x^2 + S_y^2 (1 - \rho^2) \right).$$

3.1. Efficiency Comparisons

In this section, the efficiency conditions for which the proposed modified ratio estimators performs better as compared to the existing modified ratio estimators have been derived algebraically.

3.2.1. Comparisons with existing ratio estimator

The proposed ratio estimators performed better than the existing ratio estimators if and only if,

$$MSE(\hat{\bar{Y}}_{pj}) \leq MSE(\hat{\bar{Y}}_i),$$

$$\frac{(1-f)}{n} \left(R_{pj}^2 S_x^2 + S_y^2 (1 - \rho^2) \right) \leq \frac{(1-f)}{n} \left(R_i^2 S_x^2 + S_y^2 (1 - \rho^2) \right),$$

$$R_{pj}^2 S_x^2 \leq R_i^2 S_x^2,$$

$$R_{pj} \leq R_i, \tag{2}$$

where $j = 1, 2, 3$ and $i = 1, 2, \dots, 26$.

If the above condition is fulfilled, then our suggested estimators perform better as compared to the existing estimators consider in this study.

4. Numerical Illustration

For the empirical study of the proposed and existing estimators, we have used 3 natural populations. The population 1 is obtained from Singh and Chaudhary (1986) page 177 and Population 2 and Population 3 are obtained from Murthy (1967) page 228.

The characteristics of the 3 populations are given in Table 1. The values of the constants and biases of the existing and proposed estimators are specified in Tables 2 and 3, respectively, whereas the mean square errors values of the existing and proposed estimators are given in Tables 4 and 5, respectively.

Table 1: Characteristics of the Populations.

Parameters	Pop. 1	Pop. 2	Pop. 3
N	34	80	80
n	20	20	20
\bar{Y}	856.412	5182.637	5182.637
\bar{X}	199.441	285.125	1126.463
ρ	0.445	0.915	0.941
S_y	733.141	1835.659	1835.659
C_y	0.8561	0.354	0.354
S_x	150.215	270.429	845.61
C_x	0.753	0.948	0.751
β_2	1.0445	1.301	-0.063
β_1	1.1823	0.698	1.050
M_d	142.500	148.000	757.500
QD	89.375	179.375	588.125
G	162.996	279.711	904.081
D	144.481	247.938	801.381
S_{pw}	142.99	244.838	791.364
DM	206.9444	276.189	1150.700

From an analysis of Tables 2-5, several interesting observations can be made:

- It is observed that the condition mentioned in equation (2) is satisfied because all the proposed estimators have smaller values of constants as compared to the existing estimators (cf Table 2 vs Table 3).
- The biases of the suggested estimators are smaller than the existing estimators in literature (cf Table 2 vs Table 3).
- It can be seen that the proposed estimators have lesser values of MSE as compared to the estimators proposed by Kadilar and Cingi (2004), Kadilar and Cingi (2006), Subramani and Kumarapandiyan (2012a), Subramani and Kumarapandiyan (2012b), Subramani and Kumarapandiyan (2012c), Jeelani et al. (2013) and Abid et al. (2016c) which indicates that the proposed estimators are more efficient as compared to the existing estimators (cf Table 4 vs Table 5).

Table 2: The biases and constants of existing estimators

Estimators	Constant			Bias		
	Pop. 1	Pop. 2	Pop. 3	Pop. 1	Pop. 2	Pop. 3
\hat{Y}_1	4.294	18.177	4.601	10.00	174.83	109.52
\hat{Y}_2	4.294	18.177	4.601	9.93	173.67	109.37
\hat{Y}_3	4.278	18.116	4.598	9.89	173.98	109.53
\hat{Y}_4	4.272	18.132	4.601	9.93	173.18	111.86
\hat{Y}_5	4.279	18.090	4.650	9.87	173.93	109.53
\hat{Y}_6	4.264	18.130	4.601	9.96	173.71	109.34
\hat{Y}_7	4.285	18.119	4.597	9.94	173.65	109.27
\hat{Y}_8	4.281	18.115	4.596	9.83	173.57	109.36
\hat{Y}_9	4.258	18.111	4.598	9.96	173.23	112.46
\hat{Y}_{10}	4.285	18.094	4.662	9.77	173.9	109.53
\hat{Y}_{11}	4.244	18.128	4.601	9.89	173.24	109.31
\hat{Y}_{12}	4.269	18.094	4.597	9.91	174.17	109.53
\hat{Y}_{13}	4.275	18.143	4.601	3.40	75.76	39.15
\hat{Y}_{14}	2.504	11.966	2.751	2.63	73.02	30.47
\hat{Y}_{15}	2.204	11.747	2.427	3.89	89.31	40.69
\hat{Y}_{16}	2.676	12.991	2.804	3.53	57.48	1.186
\hat{Y}_{17}	2.550	10.422	0.478	5.26	79.42	48.85
\hat{Y}_{18}	2.363	9.175	2.552	3.03	44.55	33.71
\hat{Y}_{19}	2.059	8.935	2.224	2.30	42.25	25.58
\hat{Y}_{20}	1.515	8.772	2.483	1.24	40.72	31.91
\hat{Y}_{21}	1.635	9.320	2.620	1.45	45.96	35.53
\hat{Y}_{22}	2.189	9.483	2.362	2.60	47.58	28.87
\hat{Y}_{23}	2.490	9.722	2.688	3.36	50.02	37.39
\hat{Y}_{24}	1.645	9.377	2.635	1.47	46.53	35.91
\hat{Y}_{25}	2.200	9.540	2.377	2.63	48.16	29.22
\hat{Y}_{26}	2.501	9.779	2.702	3.39	50.61	37.78

Table 3: The biases and constants of proposed estimators

Estimators	Constant			Bias		
	Pop. 1	Pop. 2	Pop. 3	Pop. 1	Pop. 2	Pop. 3
\hat{Y}_{p1}	2.107	9.233	2.276	2.137	45.111	26.800
\hat{Y}_{p2}	1.806	8.993	1.949	1.483	42.792	19.650
\hat{Y}_{p3}	1.289	8.829	2.206	0.800	41.252	25.188

Table 4: Mean square errors of existing estimators

Estimators	Mean Square Error		
	Pop. 1	Pop. 2	Pop. 3
\hat{Y}_1	17437.65	926660.7	581994.2
\hat{Y}_2	17373.31	920662.5	581238.5
\hat{Y}_3	17348.62	922242.5	582058.1
\hat{Y}_4	17376.04	918082.1	594119.8
\hat{Y}_5	17319.75	922003.4	582079.3
\hat{Y}_6	17399.52	920873.2	581046.8
\hat{Y}_7	17387.08	920560.3	580732.7
\hat{Y}_8	17294.19	920108.2	581191.4
\hat{Y}_9	17401.14	918382.8	597260.9
\hat{Y}_{10}	17239.66	921833.6	582062.1
\hat{Y}_{11}	17336.98	918450.9	580937.6
\hat{Y}_{12}	17362.26	923260.7	582055.1
\hat{Y}_{13}	11785.7	413230.8	217319.8
\hat{Y}_{14}	11127.47	399044.9	172323.8
\hat{Y}_{15}	12199.76	483450.4	225319.5
\hat{Y}_{16}	11892.07	318486.7	20545.47
\hat{Y}_{17}	13376.04	432164.6	267595.2
\hat{Y}_{18}	11465.43	251457.8	189080.3
\hat{Y}_{19}	10841.88	239515.6	146971.8
\hat{Y}_{20}	9937.20	236591.3	179770.5
\hat{Y}_{21}	10113.06	258768.3	198518.7
\hat{Y}_{22}	11097.24	267178.4	164020.7
\hat{Y}_{23}	11752.23	279802.0	208187.2
\hat{Y}_{24}	10129.01	261697.4	200516.2
\hat{Y}_{25}	11119.84	270154.6	165857.4
\hat{Y}_{26}	11777.26	282843.6	210216.9

Table 5: Mean square errors of proposed estimators

Estimators	Mean Square Error		
	Pop. 1	Pop. 2	Pop. 3
\hat{Y}_{p1}	10934.74	254364.7	153292.6
\hat{Y}_{p2}	10386.83	242346.5	116239.3
\hat{Y}_{p3}	9644.04	234368.4	144936.7

- It can be seen that the proposed estimators have smaller values of MSE as compared to the estimators proposed by Kadilar and Cingi (2004), Kadilar and Cingi (2006), Subramani and Kumarapandiyan (2012a), Subramani and Kumarapandiyan (2012b), Subramani and Kumarapandiyan (2012c), Jeelani et al. (2013) and Abid et al. (2016c) which shows that the proposed estimators perform better as compared to the existing estimators (cf Table 4 vs Table 5).
- It is noted that the proposed estimator, \hat{Y}_{P3} has a smaller MSE value i.e. (9644.04, 234368.4 and 144936.7) as compared to all the proposed estimators and existing estimator for three real populations considered in this study (cf Table 4 vs Table 5).

So in general, we can say that our suggested modified ratio estimators perform more efficiently as compared to the existing modified ratio estimators.

5. Conclusions

In sample survey the availability of auxiliary information enhances the efficiency of the estimators. In this study, we have proposed several modified ratio estimators using known value of population decile mean, coefficient of variation and population correlation coefficient by using the information on the auxiliary variable. It is observed that the mean squared error values of the suggested estimators are smaller than existing estimators for all the selected three known populations. Also, as we know that the measures like the mean, median and mode are affected by the extreme values in the population, while decile mean is robust to extreme values. Hence, it can be claimed that our proposed estimators outperform the existing estimators for the practical consideration.

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